### The Mean-Field Ensemble Kalman Filter

#### Andrew Stuart

Computing and Mathematical Sciences California Institute of Technology

#### DoD, NSF, ONR

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Leçons Jacques-Louis Lions Lectures Sorbonne Université, Paris

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### Collaborators

Edo Calvello (Caltech)Sebastian Reich (Potsdam)

Reference: E. Calvello, S. Reich, and S

Ensemble Kalman Methods: a Mean-Field Perspective. arXiv:2209.11371 (2022).

José Carrillo (Oxford)

- Franca Hoffmann (Caltech)
- Urbain Vaes (CERMICS)

Reference: J. A. Carrillo, F. Hoffmann, S, and U. Vaes

The Ensemble Kalman Filter in the Near-Gaussian Setting. arXiv: 2212.13239 (2022).

### Overview

Mean-Field Optimization Perspective

Probabilistic Perspective: True Filter

Probabilistic Perspective: Ensemble Kalman Filter

Main Theorem: Relating The True and Ensemble Kalman Filters

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Closing

# Mean-Field Optimization Perspective

Optimization: Albers, Blancquart, Levine, Seylabi and S [1] (2022)

Mean-Field: Calvello, Reich and S [3] (2022)

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# Kalman Filter (Navigation)

### State Space Model

Dynamics Model:  $v_{n+1} = Mv_n + \xi_n$ ,  $n \in \mathbb{Z}^+$ Data Model:  $y_{n+1} = Hv_{n+1} + \eta_{n+1}$ ,  $n \in \mathbb{Z}^+$ Probabilistic Structure:  $v_0 \sim N(m_0, C_0)$ ,  $\xi_n \sim N(0, \Sigma)$ ,  $\eta_n \sim N(0, \Gamma)$ Probabilistic Structure:  $v_0 \perp \{\xi_n\} \perp \{\eta_n\}$  independent



- Rudolph Kalman [19] (1960).
- ▶  $\approx$  43,500 citations (Google Scholar 12/23).

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- Apollo 11.
- The Algorithm:
- $\triangleright Y_n^{\dagger} = \{y_{\ell}^{\dagger}\}_{\ell=1}^n.$
- ►  $v_n^{\dagger} | Y_n^{\dagger} \sim \mathrm{N}(m_n, C_n).$
- $\blacktriangleright (m_n, C_n) \mapsto (m_{n+1}, C_{n+1}).$

# Kalman Filter

### Sequential Optimization Viewpoint

$$\begin{array}{ll} {\sf Predict:} & \widehat{m}_{n+1} = Mm_n, & n \in \mathbb{Z}^+ \\ {\sf Model/Data \ Compromise:} & J_n(m) = \frac{1}{2} |m - \widehat{m}_{n+1}|^2_{\widehat{C}_{n+1}} + \frac{1}{2} |y_{n+1}^\dagger - Hm|^2_{\Gamma} \\ {\sf Optimize:} & m_{n+1} = \operatorname{argmin}_m J_n(m). \end{array}$$

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► | · | Euclidean norm on any R<sup>r</sup> and induced matrix norm.

• 
$$|\cdot|_A = |A^{-\frac{1}{2}} \cdot |$$
 for  $A > 0$  spd.

- *d* the state space dimension  $(m_n, v_n \in \mathbb{R}^d)$ .
- Updating  $\widehat{C}_{n+1}$  is expensive:  $\mathcal{O}(d^2)$  storage.

# 3DVAR Filter (Weather Forecasting)

### State Space Model

Dynamics Model:  $v_{n+1} = \Psi(v_n) + \xi_n$ ,  $n \in \mathbb{Z}^+$ Data Model:  $y_{n+1} = Hv_{n+1} + \eta_{n+1}$ ,  $n \in \mathbb{Z}^+$ Probabilistic Structure:  $v_0 \sim N(m_0, C_0)$ ,  $\xi_n \sim N(0, \Sigma)$ ,  $\eta_n \sim N(0, \Gamma)$ Probabilistic Structure:  $v_0 \perp \{\xi_n\} \perp \{\eta_n\}$  independent



- Andrew Lorenc [23] (1986).
- $\approx$  2,000 citations (Google Scholar 12/23).
- Introduced in UK Met Office.
- The Algorithm:
- $\blacktriangleright \{v_n\} \mapsto \{v_{n+1}\}.$
- Given  $Y_n^{\dagger}$  want  $v_n \approx v_n^{\dagger}$ . (Again  $Y_n^{\dagger} = \{y_{\ell}^{\dagger}\}_{\ell=1}^n$ .)

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# 3DVAR

### Sequential Optimization Viewpoint

$$\begin{array}{ll} \text{Predict:} & \widehat{v}_{n+1} = \Psi(v_n), & n \in \mathbb{Z}^+\\ \text{Model/Data Compromise:} & J_n(v) = \frac{1}{2} |v - \widehat{v}_{n+1}|_{\widehat{C}}^2 + \frac{1}{2} |y_{n+1}^{\dagger} - Hv|_{\widehat{r}}^2\\ \text{Optimize:} & v_{n+1} = \operatorname{argmin}_v J_n(v). \end{array}$$

- $\hat{C}$  is a fixed model covariance (not updated sequentially).
- $d = \mathcal{O}(10^9)$ ;  $\mathcal{O}(d^2)$  entries of  $\hat{C}$  prohibitive in general.
- $\hat{C}$  chosen to have simple, computable, structure.

## Ensemble Kalman Filter

#### State Space Model

Dynamics Model:  $v_{n+1} = \Psi(v_n) + \xi_n$ ,  $n \in \mathbb{Z}^+$ Data Model:  $y_{n+1} = Hv_{n+1} + \eta_{n+1}$ ,  $n \in \mathbb{Z}^+$ Probabilistic Structure:  $v_0 \sim N(m_0, C_0)$ ,  $\xi_n \sim N(0, \Sigma)$ ,  $\eta_n \sim N(0, \Gamma)$ Probabilistic Structure:  $v_0 \perp \{\xi_n\} \perp \{\eta_n\}$  independent



- Geir Evensen [11] (1994).
- ▶  $\approx$  6,000 citations (Google Scholar 12/23).
- Originally ocean dynamics; now weather.
- $\mu_n := \operatorname{Law}(\mathbf{v}_n^{\dagger} | \mathbf{Y}_n^{\dagger})$ . (Here  $\mathbf{Y}_n^{\dagger} = \{\mathbf{y}_{\ell}^{\dagger}\}_{\ell=1}^n$ .)
- Mean-Field Algorithm:
- $\blacktriangleright (v_n, \mu_n^{\mathsf{EK}}) \mapsto (v_{n+1}, \mu_{n+1}^{\mathsf{EK}}). \quad \mu_n^{\mathsf{EK}} := \operatorname{Law}(v_n).$
- When is this approximation valid:  $\mu_n^{EK} \approx \mu_n$ ?

# Mean-Field Ensemble Kalman Filter

### Sequential Optimization Viewpoint

$$\begin{array}{ll} \mathsf{Predict:} & \widehat{v}_{n+1} = \Psi(v_n) + \xi_n, \quad n \in \mathbb{Z}^+\\ \mathsf{Model/Data \ Compromise:} & J_n(v) = \frac{1}{2} |v - \widehat{v}_{n+1}|_{\widehat{\mathcal{C}}_{n+1}}^2 + \frac{1}{2} |y_{n+1}^{\dagger} + \eta_{n+1} - Hv|_{\Gamma}^2\\ \mathsf{Optimize:} & v_{n+1} = \operatorname{argmin}_v \ J_n(v). \end{array}$$

### In What Sense Is This A Mean-Field Model?

$$\blacktriangleright \ \mu_n^{EK} := \operatorname{Law}(v_n).$$

- $\widehat{C}_{n+1}$  is the covariance under  $\widehat{\mu}_{n+1}^{EK} := \operatorname{Law}(\widehat{v}_{n+1}).$
- ▶  $|\cdot|$  Euclidean norm on **R**<sup>*r*</sup> and induced matrix norm (any *r*).

• 
$$|\cdot|_{A} = |A^{-\frac{1}{2}} \cdot |$$
 for  $A > 0$  spd

# Mean-Field Ensemble Kalman Filter

### Sequential Optimization Viewpoint

$$\begin{array}{ll} {\sf Predict:} & \widehat{v}_{n+1} = \Psi(v_n) + \xi_n, \quad n \in \mathbb{Z}^+ \\ {\sf Model/Data \ Compromise:} & J_n(v) = \frac{1}{2} |v - \widehat{v}_{n+1}|_{\widehat{\mathcal{C}}_{n+1}}^2 + \frac{1}{2} |y_{n+1}^{\dagger} + \eta_{n+1} - Hv|_{\Gamma}^2 \\ {\sf Optimize:} & v_{n+1} = \operatorname{argmin}_v \ J_n(v). \end{array}$$

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• 
$$|\cdot|_A = |A^{-\frac{1}{2}} \cdot |$$
 for  $A > 0$  spd

- ▶ In practice: use  $j \in \{1, ..., J\}$ , J number of ensemble members.
- Use resulting ensemble  $\hat{v}_{n+1}^{(j)}$  to estimate  $\hat{C}_{n+1}$ .

# Summary Of Optimization Perspective



Two Goals

Control (3DVAR, EnKF):  $|v_n - v_n^{\dagger}| \ll 1$ , next slide. UQ (EnKF):  $\mu_n^{EK} \approx \mu_n = \text{Law}(v_n^{\dagger}|Y_n^{\dagger})$ , rest of talk.

# 3DVAR and Small Noise

Synchronization and Lorenz '63 Pecora and Carroll [24] (1990)

Synchronization and Navier-Stokes Hayden, Olson and Titi [15] (2011)

#### Theorem Law, Shukla and S [21] (2012)

Assume synchronization and small noise  $\mathcal{O}(\epsilon)$  in truth, no noise in filter. Consider 3DVAR with  $K = \gamma H^*$  and  $|\gamma - 1| \leq 1$ . Then

$$\limsup_{n\to\infty}\mathbb{E}\Big|v_n-v_n^{\dagger}\Big|^2\leq C\epsilon^2.$$

Corollary Sanz-Alonso and S [29] (2015)

Assume synchronization and small noise  $\mathcal{O}(\epsilon)$  in truth, no noise in filter. The true filtering distribution  $\mu_n = \text{Law}(v_n^{\dagger}|Y_n^{\dagger})$  satisfies

$$\limsup_{n\to\infty}\mathbb{E}\left|\mathbb{E}^{\nu\sim\mu_n}\nu-\nu_n^{\dagger}\right|^2\leq C\epsilon^2.$$

# Probabilistic Perspective: True Filter

Filtering: Doucet, de Freitas and Gordon [10] (2004)

Vectors and Matrices

Measures and Densities

Operators

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## Unconditioned Dynamics

### State-Data Viewpoint (Nonlinear)

State:
 
$$v_{n+1} = \Psi(v_n) + \xi_n,$$
 $\xi_n \sim N(0, \Sigma), \text{ i.i.d. },$ 

 Data:
  $y_{n+1} = h(v_{n+1}) + \eta_{n+1},$ 
 $\eta_{n+1} \sim N(0, \Gamma), \text{ i.i.d. },$ 
 $v_0 \sim N(m_0, C_0),$ 
 $v_0 \perp \{\xi_n\}_{n \in \mathbb{N}} \perp \{\eta_{n+1}\}_{n \in \mathbb{N}}$ 

Probability Viewpoint (Linear)

$$v_n \sim \pi_n, \quad (v_n, y_n) \sim \mathfrak{r}_n,$$
  
$$\pi_{n+1} = P \pi_n,$$
  
$$\mathfrak{r}_{n+1} = Q \pi_{n+1}$$

# Key Linear Operators on $\mathcal P$

### Definition of $\mathcal{P}$ , $\mathcal{G}$

- $\mathcal{P}(\mathbf{R}^r)$  : all probability measures on  $\mathbf{R}^r$ .
- $\mathcal{G}(\mathbf{R}^r)$ : all Gaussian probability measures on  $\mathbf{R}^r$ .

### Definition of P

 $P: \mathcal{P}(\mathbf{R}^d) \to \mathcal{P}(\mathbf{R}^d)$  is the linear operator:

$$\boldsymbol{P}\pi(\boldsymbol{u}) = \frac{1}{\sqrt{(2\pi)^d \det \Sigma}} \int \exp\left(-\frac{1}{2}|\boldsymbol{u} - \Psi(\boldsymbol{v})|_{\Sigma}^2\right) \pi(\boldsymbol{v}) \,\mathrm{d}\boldsymbol{v}.$$

### Definition of Q

 $Q \colon \mathcal{P}(\mathbf{R}^d) \to \mathcal{P}(\mathbf{R}^d \times \mathbf{R}^K)$  is the linear operator:

$$Q\pi(u,y) = rac{1}{\sqrt{(2\pi)^{K}\det\Gamma}}\exp\left(-rac{1}{2}|y-h(u)|_{\Gamma}^{2}
ight)\pi(u).$$

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# Key Nonlinear Operator on ${\mathcal P}$



#### Conditioning (Nonlinear)

 $B(\bullet; y^{\dagger}): \mathcal{P}(\mathbf{R}^{d} \times \mathbf{R}^{K}) \rightarrow \mathcal{P}(\mathbf{R}^{d})$  describes conditioning on observation  $y = y^{\dagger}$ :

$$B(\rho; y^{\dagger})(u) = \frac{\rho(u, y^{\dagger})}{\int_{\mathbf{R}^d} \rho(u, y^{\dagger}) \, \mathrm{d}u}$$

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# The True Filter

#### Sequential Interleaving of Prediction and Bayes Theorem

 $P\mu_n$  is prior prediction;  $L(\bullet; y^{\dagger}) := B(\bullet; y^{\dagger}) \circ Q$  maps prior to posterior:

 $\mu_{n+1} = B(QP\mu_n; y_{n+1}^{\dagger}),$  $\mu_{n+1} = L(P\mu_n; y_{n+1}^{\dagger}).$ 

### Particle Filter Doucet [10] (2015)

 $S^{J}: \mathcal{P}(\mathbf{R}^{r}) \times \Omega \rightarrow \mathcal{P}(\mathbf{R}^{r})$  is empirical approximation operator:

$$S^{J}\mu = \frac{1}{J}\sum_{j=1}^{J}\delta_{\mathbf{v}_{j}}, \quad \mathbf{v}_{j}\sim\mu \text{ i.i.d.}.$$

 $S^{J}$ : is thus a random approximation of the identity operator on  $\mathcal{P}(\mathbf{R}^{r})$ .

$$\mu_{n+1}^{PF} = L(S^J P \mu_n^{PF}; y_{n+1}^{\dagger}).$$

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# Particle Filter Convergence

Theorem Del Moral [7] (1997), Del Moral and Guionnet [9] (2001)

$$\sup_{0\leq n\leq N}d(\mu_n,\mu_n^{PF})\leq \frac{C}{\sqrt{J}}.$$

Comments on Proof Rebschini and Van Handel [25] (2015),

Metric d(·, ·) on random probability measures:

• 
$$d(\mu, \nu)^2 = \sup_{|f| \le 1} \mathbb{E} |\mu(f) - \nu(f)|^2$$

- Reduces to TV between deterministic measures.
- Consistency + Stability Implies Convergence.
- Consistency:  $d(S^{J}\mu, \mu) \leq \frac{1}{\sqrt{J}}$ .
- **Stability**: P, L Lipschitz in  $d(\cdot, \cdot)$ .
- Suffers from weight collapse.

# Weights

### Particle Filter (Weight Collapse)

$$\begin{split} \widehat{v}_{n+1}^{(j)} &= \Psi\left(v_n^{(j)}\right) + \xi_n^{(j)}, \quad v_n^{(j)} \sim \mu_n^{PF}, \\ \ell_{n+1}^{(j)} &= \exp\left(-\frac{1}{2} \big| y_{n+1}^{\dagger} - h\big(\widehat{v}_{n+1}^{(j)}\big)\big|_{\Gamma}^2\right), \\ \mu_{n+1}^{PF} &= \sum_{j=1}^J w_{n+1}^{(j)} \delta_{\widehat{v}_{n+1}^{(j)}}, \quad w_{n+1}^{(j)} = \ell_{n+1}^{(j)} \Big/ \big(\sum_{m=1}^J \ell_{n+1}^{(m)}\big) \end{split}$$

Ensemble Kalman Filter (No Weight Collapse!)

$$\begin{split} \widehat{v}_{n+1}^{(j)} &= \Psi \big( v_n^{(j)} \big) + \xi_n^{(j)}, \quad v_n^{(j)} \sim \mu_n^{EK}, \\ \widehat{y}_{n+1}^{(j)} &= h \big( \widehat{v}_{n+1}^{(j)} \big) + \eta_{n+1}^{(j)}, \\ v_{n+1}^{(j)} &= \widehat{v}_{n+1}^{(j)} + \mathcal{C}^{vy} \big( \rho_{n+1}^{EK,J} \big) \mathcal{C}^{yy} \big( \rho_{n+1}^{EK,J} \big)^{-1} \big( y_{n+1}^{\dagger} - \widehat{y}_{n+1}^{(j)} \big), \\ \mu_{n+1}^{EK} &= \frac{1}{J} \sum_{j=1}^J \delta_{v_{n+1}^{(j)}}. \end{split}$$

# Probabilistic Perspective: Ensemble Kalman Filter

Mean-Field: Calvello, Reich and S [3] (2022)

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## The Mean Field Ensemble Kalman Filter

Comparison With True Filter

$$\mu_{n+1}^{EK} = T(QP\mu_n^{EK}; y_{n+1}^{\dagger}),$$
  
$$\mu_{n+1} = B(QP\mu_n; y_{n+1}^{\dagger}).$$

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#### Observations About *T*

- Choose T to recover mean-field EnKF;
- T defined through pushforward;
- Key is to understand when  $T \approx B$ .
- $\blacktriangleright T \equiv B \text{ on } \mathcal{G}(\mathbb{R}^d \times \mathbb{R}^K).$

# Approximate Conditioning

Block Form Of State-Data Covariance Write covariance under  $\rho \in \mathcal{P}(\mathbf{R}^d \times \mathbf{R}^{\kappa})$  as:

$$\operatorname{cov}_{\rho} = \begin{pmatrix} \mathcal{C}^{\mathsf{vv}}(\rho) & \mathcal{C}^{\mathsf{vy}}(\rho) \\ \mathcal{C}^{\mathsf{vy}}(\rho)^{\mathsf{T}} & \mathcal{C}^{\mathsf{yy}}(\rho) \end{pmatrix}.$$

Key Nonlinear Operator on  $\mathcal{P}$   $(\tau(\bullet; y^{\dagger}) \equiv B(\bullet; y^{\dagger})$  for Gaussian inputs)  $\mathcal{T}(\bullet; y^{\dagger}) \colon \mathcal{P}(\mathbf{R}^{d} \times \mathbf{R}^{K}) \to \mathcal{P}(\mathbf{R}^{d})$  approximates conditioning of  $\rho$  on  $y = y^{\dagger}$ :  $\mathfrak{T}(\bullet, \bullet; \rho, y^{\dagger}) \colon \mathbf{R}^{d} \times \mathbf{R}^{K} \to \mathbf{R}^{d};$   $(v, y) \mapsto v + \mathcal{C}^{vy}(\rho)\mathcal{C}^{yy}(\rho)^{-1}(y^{\dagger} - y),$   $\mathcal{T}(\rho; y^{\dagger}) = (\mathfrak{T}(\bullet, \bullet; \rho, y^{\dagger}))_{\sharp}\rho,$  $\mu_{n+1}^{EK} = \mathcal{T}(\mathcal{QP}\mu_{n}^{EK}; y_{n+1}^{\dagger}).$ 

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# Mean Field EnKF & Maps on Probability Measures

State-Data Space Picture

$$\begin{split} \widehat{v}_{n+1} &= \Psi(v_n) + \xi_n, \\ \widehat{y}_{n+1} &= h(\widehat{v}_{n+1}) + \eta_{n+1}, \\ v_{n+1} &= \widehat{v}_{n+1} + \mathcal{C}^{vy}(\rho_{n+1}^{EK}) \mathcal{C}^{yy}(\rho_{n+1}^{EK})^{-1}(y_{n+1}^{\dagger} - \widehat{y}_{n+1}). \end{split}$$

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### Remarks

- Recovers mean-field EnKF in nudging form.
- Use equal weight particle approximation to implement.

# Main Theorem Relating The True and Ensemble Kalman Filters

Main Theorem: Carrillo, Hoffmann, S and Vaes [4] (2022)

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# The Mean Field Ensemble Kalman Filter

#### Comparison With True Filter

$$\mu_{n+1}^{EK} = T(QP\mu_n^{EK}; y_{n+1}^{\dagger}),$$
  
$$\mu_{n+1} = B(QP\mu_n; y_{n+1}^{\dagger}),$$

#### Remarks

- When is  $T(\bullet; y^{\dagger}) \approx B(\bullet; y^{\dagger})$ ?
- A form of Consistency.
- Try Consistency + Stability Implies Convergence.

### Gaussian Projection

#### Best Gaussian Approximation in KL

$$G: \mathcal{P} \to \mathcal{G},$$
  
$$G\pi = \operatorname{argmin}_{\mathfrak{p} \in \mathcal{G}} d_{\mathrm{KL}}(\pi \| \mathfrak{p}).$$

#### Best Gaussian Approximation in KL

 $G\pi = N(\text{mean}_{\pi}, \text{cov}_{\pi}).$ 

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### The Mean Field Ensemble Kalman Filter

#### Comparison With True Filter

$$\mu_{n+1}^{EK} = T(QP\mu_n^{EK}; y_{n+1}^{\dagger}),$$
$$\mu_{n+1} = B(QP\mu_n; y_{n+1}^{\dagger}).$$

#### Key Fact

$$T(G\rho; y^{\dagger}) = B(G\rho; y^{\dagger}) \quad \forall (\rho, y^{\dagger}) \in \mathcal{P}(\mathbf{R}^{d} \times \mathbf{R}^{EK}) \times \mathbf{R}^{EK}$$

Pushforward beyond the Gaussian setting (continuous time): Yang, Mehta and Meyn [33] (2013) Pushforward beyond the Gaussian setting (discrete time): Spantini, Baptista and Marzouk [32] (2022)

# Closness of Exact Filter and EnKF

### Weighted TV Metric

Let 
$$g(v) = 1 + |v|^2$$
.  
 $d_g(\mu_1, \mu_2) = \sup_{|f| \le g} |\mu_1[f] - \mu_2[f]|, \quad \mu[f] = \int f(u)\mu(du)$ 

### Definition

Measure of how close true filter  $\{\mu_n\}$  is to being Gaussian:

 $\varepsilon := \sup_{0 \le n \le N} d_g(GQP\mu_n, QP\mu_n).$ 

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# Closness of Exact Filter and EnKF

### Weighted TV Metric

Let 
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#### Definition

Measure of how close true filter  $\{\mu_n\}$  is to being Gaussian:

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Theorem Carrillo, Hoffmann, S and Vaes [4] (2022)

Let  $\mu_0^{EK} = \mu_0$  and assume that  $\|\Psi\|_{L^{\infty}}, \|h\|_{L^{\infty}}$  and  $|h|_{C^{0,1}}$  are bounded by r. Then there is C := C(N, r) > 0:

$$\sup_{0\leq n\leq N}d_g(\mu_n,\mu_n^{\mathsf{EK}})\leq C\varepsilon.$$

# Closness of Exact Filter and EnKF

# Assumptions C • Data $Y_j^{\dagger}$ lies in set $B_y := \left\{ Y^{\dagger} \in \mathbf{R}^{KJ} : \max_{0 \le j \le J} |y_j^{\dagger}| \le \kappa_y \right\}.$

▶  $\Psi_0$ :  $\mathbf{R}^d \to \mathbf{R}^d$  and  $\mathbf{h}_0$ :  $\mathbf{R}^d \to \mathbf{R}^K$  are constant functions and denote by  $B_{\Psi,h}(r)$  the set  $(\Psi, h)$  satisfying  $\Psi \in B_{L^{\infty}}(\Psi_0, r)$ ,  $h \in B_{L^{\infty}}(h_0, r)$ .

Corollary Carrillo, Hoffmann, S and Vaes [4] (2022)

Let Assumptions T of Theorem hold and Assumptions C. Then for any  $\epsilon>0$  there is  $\delta>0$  such that

$$\sup_{Y^{\dagger}\in B_{Y}}\sup_{(\Psi,h)\in B_{\Psi,h}(\delta)}\sup_{0\leq n\leq N}d_{g}(\mu_{n},\mu_{n}^{EK})\leq\epsilon.$$

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# Proof of Theorem (Lipschitz Estimates)

### Linear Maps P, Q

The maps P, Q are globally Lipschitz on  $\mathcal{P}(\mathbf{R}^d)$  in  $d_g$ .



# Proof of Theorem (Stability Estimate I)

Conditioning is not Lipschitz stable. However, if  $\Psi$  is bounded:

Nonlinear Conditioning Map  $B^{y^{\dagger}}$ The maps  $B^{y^{\dagger}}(\bullet) := B(\bullet; y^{\dagger})$  satisfy:  $\forall \mu \in \mathcal{P}(\mathbf{R}^{d})$  $d_{g}(B^{y^{\dagger}}(GQP\mu), B^{y^{\dagger}}(QP\mu)) \leq \ell_{B} d_{g}(GQP\mu, QP\mu).$ 

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# Proof of Theorem (Stability Estimate II)

Let  $\mathcal{P}_R$  denote the following subset of probability measures

$$\mathcal{P}_{R}(\mathbf{R}^{r}) = \left\{ \mu \in \mathcal{P}(\mathbf{R}^{r}) : \max\left\{ |\operatorname{mean}(\mu)|, |\operatorname{cov}(\mu)|^{\frac{1}{2}}, |\operatorname{cov}(\mu)|^{-\frac{1}{2}} \right\} \le R \right\}.$$

Using linearity of  $\mathfrak{T}$ , which defines nonlinear map  $T^{y^{\dagger}}$ :

Approximate Nonlinear Conditioning Map  $T^{y^{\dagger}}$ The maps  $T^{y^{\dagger}}(\bullet) := T(\bullet; y^{\dagger})$  satisfy, using  $\Psi$  bounded,  $\forall (\mu, \rho) \in \mathcal{P}(\mathbf{R}^{d}) \times \mathcal{P}_{R}(\mathbf{R}^{d} \times \mathbf{R}^{K}),$  $d_{g}(T^{y^{\dagger}}(QP\mu), T^{y^{\dagger}}(\rho)) \leq \ell_{T}(R) d_{g}(QP\mu, \rho),$ 

# Proof of Theorem

Since 
$$T^{y_{n+1}^{\dagger}}(G_{\bullet}) = B^{y_{n+1}^{\dagger}}(G_{\bullet})$$
 we have  
 $d_g(\mu_{n+1}^{EK}, \mu_{n+1}) = d_g\left(T^{y_{n+1}^{\dagger}}(QP\mu_n^{EK}), B^{y_{n+1}^{\dagger}}(QP\mu_n)\right)$   
 $\leq d_g\left(T^{y_{n+1}^{\dagger}}(QP\mu_n^{EK}), T^{y_{n+1}^{\dagger}}(QP\mu_n)\right)$   
 $+ d_g\left(T^{y_{n+1}^{\dagger}}(QP\mu_n), T^{y_{n+1}^{\dagger}}(GQP\mu_n)\right)$   
 $+ d_g\left(T^{y_{n+1}^{\dagger}}(GQP\mu_n), B^{y_{n+1}^{\dagger}}(QP\mu_n)\right)$   
 $\leq \ell_T(R) d_g\left(QP\mu_n^{EK}, QP\mu_n\right)$   
 $+ \ell_T(R) d_g\left(QP\mu_n, GQP\mu_n\right)$   
 $+ d_g\left(B^{y_{n+1}^{\dagger}}(GQP\mu_n), B^{y_{n+1}^{\dagger}}(QP\mu_n)\right)$   
 $\leq cd_g(\mu_n^{EK}, \mu_n) + (\ell_T(R) + \ell_B)\varepsilon.$ 

# Proof of Theorem

Since 
$$T^{y_{n+1}^{\dagger}}(G_{\bullet}) = B^{y_{n+1}^{\dagger}}(G_{\bullet})$$
 we have  
 $d_g(\mu_{n+1}^{EK}, \mu_{n+1}) = d_g\left(T^{y_{n+1}^{\dagger}}(QP\mu_n^{EK}), B^{y_{n+1}^{\dagger}}(QP\mu_n)\right)$   
 $\leq d_g\left(T^{y_{n+1}^{\dagger}}(QP\mu_n^{EK}), T^{y_{n+1}^{\dagger}}(QP\mu_n)\right)$   
 $+ d_g\left(T^{y_{n+1}^{\dagger}}(QP\mu_n), T^{y_{n+1}^{\dagger}}(GQP\mu_n)\right)$   
 $+ d_g\left(T^{y_{n+1}^{\dagger}}(GQP\mu_n), B^{y_{n+1}^{\dagger}}(QP\mu_n)\right)$   
 $\leq \ell_T(R) d_g\left(QP\mu_n^{EK}, QP\mu_n\right)$   
 $+ \ell_T(R) d_g\left(QP\mu_n, GQP\mu_n\right)$   
 $+ d_g\left(B^{y_{n+1}^{\dagger}}(GQP\mu_n), B^{y_{n+1}^{\dagger}}(QP\mu_n)\right)$   
 $\leq cd_g(\mu_n^{EK}, \mu_n) + (\ell_T(R) + \ell_B)\varepsilon.$ 

# Closing

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# Conclusions

- Introduced in 1960 by Rudolph Kalman.
- Basic algorithm generalized: 3DVAR, Ensemble Kalman (EK).
- EK methods:
  - developing as a general methodology for state estimation;
  - developing as a general methodology for inverse problems.
- EK methods applied in numerous fields:
  - weather forecasting;
  - oceanography;
  - hydrology, subsurface flow;
  - medical imaging, machine learning · · · .
- Analysis in its infancy:
  - accuracy of 3DVAR (State Estimation) last decade.
  - accuracy of EK (UQ) end of last year.
- Many open mathematical questions: great field to enter!

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